

- 가 가 가 -

Determining the value of reductions in radiation risk using the contingent valuation method: Results for the public

,

19

가	가	가			
		가	.	.	
		,		20%	
가		.			가
60.8	,			51.9	.
		,			.

Abstract

In the authors' previous study an Internet survey had been conducted for employees in nuclear related industries and institutes in Korea and the monetary value placed on reductions in risks from occupational radiation exposure and car accidents had been estimated. This paper presents the results of the follow-up study using the same methodology for the general public. The double bounded dichotomous choice(DBDC) approach was used and willingness-to-pay(WTP) values for predefined 20% reductions of the car accidents and radiation exposure risks were elicited. WTP for the reductions of the risk were estimated and values of a statistical life were calculated from the mean WTPs. For the general public, the mean value of statistical life was 2.67 billion Won for car accidents and 2.28 billion Won for radiation exposure. Discussions on the use of Contingent Valuation Method(CVM) for determining the monetary value of risk reductions and recommendations for further study were given.

1.

(Cost Benefit Analysis)

가

가 가 가 (Contingent

Valuation Method: CVM)

[1].

가

24.1

39.3

203 /man.rem

(USNRC)가

2000\$/man.rem

가

CVM

1 IPC

가

가

2000 9 가 가

가

1582

(+ 2.5%

95%).

2. CVM

1)

가

(Maximum Willingness-To-Pay)

20%

가

20%

가

가

가

가 GW-year 0.006-0.21

1)

“

2,500 I”

20%

1

20%

4

가

2,500 1

2) 2

(WTP)

2

(Double-bounded Dichotomous Choice : DBDC)

. DBDC

$i (i =$

1, 2, ..., N) 가

가 1

B_i

가

‘ ,

2

B_{iH}

, ‘ ,

2

B_{iL}

i

1. $i \in NN$ if and only if the response is “no-no” ($WTP_i \leq B_{iL}$)
2. $i \in NY$ if and only if the response is “no-yes” ($B_{iL} \leq WTP_i \leq B_i$)
3. $i \in YN$ if and only if the response is “yes-no” ($B_i \leq WTP_i \leq B_{iH}$)
4. $i \in YY$ if and only if the response is “yes-yes” ($B_{iH} \leq WTP_i$)

DBDC

, “no-no”

“yes-yes”

가

18.9%,

18.8%

WTP

가 . Bishop and Heberlein (1979, [2])

Cameron (1988. [3])

zero response

point mass

가

1

WTP

Krström (1995, [4])가

Spike Model

An and

Ayala (1995, [5])가

Mixture Model

Mixture Model

1)

90,000

1

2,500

1

i)

$$P_{i \in ZR} = \frac{\exp(Z_i B_1)}{1 + \exp(Z_i B_1)} \tag{1}$$

$$Z_i = (1, x_{i1}, x_{i2}, \dots, x_{ip}) \quad i$$

$$B_1 = (b_{1INT}, b_{11}, b_{12}, \dots, b_{1p})^T$$

ii)

가

$$\begin{cases} P_{i \in NN} = (1 - P_{i \in ZR}) P\{WTP_i \leq B_{iL}\} \\ P_{i \in NY} = (1 - P_{i \in ZR}) P\{B_{iL} \leq WTP_i \leq B_i\} \\ P_{i \in YN} = (1 - P_{i \in ZR}) P\{B_i \leq WTP_i \leq B_{iH}\} \\ P_{i \in YY} = (1 - P_{i \in ZR}) P\{WTP_i \geq B_{iH}\} \end{cases} \tag{2}$$

Mixture model

i)

가

가

, ii)

ii)

WTP_i

Bishop

and Heberlein

가

(3)

(4)

, Cameron

가

(5)

(6)

$$\begin{cases} u_{i0} = \alpha_0 y_i + \delta_i \\ u_{i1} = \alpha_1 y_i + \delta_i + \exp\left(\frac{X_i B_2 - \varepsilon_i}{\beta}\right) \end{cases} \tag{3}$$

$$WTP_i = \alpha y_i + \frac{1}{\alpha_1} \exp\left(\frac{X_i B_2 - \varepsilon_i}{\beta}\right), \quad \alpha = \frac{\alpha_1 - \alpha_0}{\alpha_1} \tag{4}$$

$$\begin{cases} u_{i0} = \alpha_0 y_i + \varepsilon_{i0} \\ u_{i1} = \alpha_1 y_i + (X_i B_2) + \varepsilon_{i1} \end{cases} \tag{5}$$

$$WTP_i = \frac{(\alpha_1 - \alpha_0)}{\alpha_1} y_i + \frac{1}{\alpha_1} X_i B_2 + \frac{1}{\alpha_1} \eta_i, \quad \eta_i = (\varepsilon_{i0} - \varepsilon_{i1}) \tag{6}$$

u_{i0} i , u_{i1} i , y_i i , $X_i = (1, x_{i1}, x_{i2}, \dots, x_{iq})$ i , δ_i ε , $\alpha_0, \alpha_1, \beta, \gamma, B_2 = (b_{2INT}, b_{21}, b_{22}, \dots, b_{2q})^T$ ε (standard

extreme value distribution)

(4)

WTP_i

(Weibull distribution)

(4) WTP_i

(Logistic distribution)

A

(7)

(8)

$$P\{WTP_i \leq A\} = 1 - \exp[-e^{\beta \log \alpha_1 - X_i B_2} (A - \alpha y_i)^\beta], \quad 0 \leq \alpha \leq (A/y_i)_{\max} \quad (7)$$

$$P\{WTP_i \leq A\} = \frac{1}{1 + \exp[X_i B_2 + (\alpha_1 - \alpha_0) y_i - \alpha_1 A]} \quad (8)$$

Bishop and Heberlein

zero-response

i

(9)

(10)

$$\left\{ \begin{array}{l} P_{i \in ZR} = e^{Z_i B_1} / (1 + e^{Z_i B_1}) \\ P_{i \in NN} = (1 - P_{i \in ZR}) [1 - \exp[-e^{\beta \log \alpha_1 - X_i B_2} (B_{iL} - \alpha y_i)^\beta]] \\ P_{i \in NY} = (1 - P_{i \in ZR}) [\exp[-e^{\beta \log \alpha_1 - X_i B_2} (B_{iL} - \alpha y_i)^\beta] - \exp[-e^{\beta \log \alpha_1 - X_i B_2} (B_i - \alpha y_i)^\beta]] \\ P_{i \in YN} = (1 - P_{i \in ZR}) [\exp[-e^{\beta \log \alpha_1 - X_i B_2} (B_i - \alpha y_i)^\beta] - \exp[-e^{\beta \log \alpha_1 - X_i B_2} (B_{iH} - \alpha y_i)^\beta]] \\ P_{i \in YY} = (1 - P_{i \in ZR}) \exp[-e^{\beta \log \alpha_1 - X_i B_2} (B_{iH} - \alpha y_i)^\beta] \end{array} \right. \quad (9)$$

$$\begin{aligned} \log(L) &= \sum_{i \in ZR} -\log\{1 + \exp(-Z_i B_1)\} + \sum_{i \in ZR} -\log\{1 + \exp(Z_i B_1)\} \\ &+ \sum_{i \in NN} \log[1 - \exp[-e^{\beta \log \alpha_1 - X_i B_2} (B_{iL} - \alpha y_i)^\beta]] \\ &+ \sum_{i \in NY} \log[\exp[-e^{\beta \log \alpha_1 - X_i B_2} (B_{iL} - \alpha y_i)^\beta] - \exp[-e^{\beta \log \alpha_1 - X_i B_2} (B_i - \alpha y_i)^\beta]] \\ &+ \sum_{i \in YN} \log[\exp[-e^{\beta \log \alpha_1 - X_i B_2} (B_i - \alpha y_i)^\beta] - \exp[-e^{\beta \log \alpha_1 - X_i B_2} (B_{iH} - \alpha y_i)^\beta]] \\ &+ \sum_{i \in YY} \log[\exp[-e^{\beta \log \alpha_1 - X_i B_2} (B_{iH} - \alpha y_i)^\beta]] \end{aligned} \quad (10)$$

Cameron

zero-response

i

(11)

(12)

$$\left\{ \begin{array}{l} P_{i \in ZR} = e^{Z_i B_1} / (1 + e^{Z_i B_1}) \\ P_{i \in NN} = (1 - P_{i \in ZR}) \frac{1}{1 + e^{X_i B_2 + (\alpha_1 - \alpha_0) y_i - \alpha_1 B_{iL}}} \\ P_{i \in NY} = (1 - P_{i \in ZR}) \left[\frac{1}{1 + e^{X_i B_2 + (\alpha_1 - \alpha_0) y_i - \alpha_1 B_i}} - \frac{1}{1 + e^{X_i B_2 + (\alpha_1 - \alpha_0) y_i - \alpha_1 B_{iL}}} \right] \\ P_{i \in YN} = (1 - P_{i \in ZR}) \left[\frac{1}{1 + e^{X_i B_2 + (\alpha_1 - \alpha_0) y_i - \alpha_1 B_{iH}}} - \frac{1}{1 + e^{X_i B_2 + (\alpha_1 - \alpha_0) y_i - \alpha_1 B_i}} \right] \\ P_{i \in YY} = (1 - P_{i \in ZR}) \frac{1}{1 + e^{-X_i B_2 - (\alpha_1 - \alpha_0) y_i + \alpha_1 B_{iH}}} \end{array} \right. \quad (11)$$

$$\begin{aligned}
\log(L) = & \sum_{i \in ZR} -\log\{1 + \exp(-Z_i B_1)\} + \sum_{i \notin ZR} -\log\{1 + \exp(Z_i B_1)\} \\
& + \sum_{i \in NN} \log \left[\frac{1}{1 + e^{X_i B_2 + (\alpha_1 - \alpha_0) y_i - \alpha_1 B_{iL}}} \right] \\
& + \sum_{i \in NY} \log \left[\frac{1}{1 + e^{X_i B_2 + (\alpha_1 - \alpha_0) y_i - \alpha_1 B_i}} - \frac{1}{1 + e^{X_i B_2 + (\alpha_1 - \alpha_0) y_i - \alpha_1 B_{iL}}} \right] \quad (12) \\
& + \sum_{i \in YN} \log \left[\frac{1}{1 + e^{X_i B_2 + (\alpha_1 - \alpha_0) y_i - \alpha_1 B_{iH}}} - \frac{1}{1 + e^{X_i B_2 + (\alpha_1 - \alpha_0) y_i - \alpha_1 B_i}} \right] \\
& + \sum_{i \in YY} \log \left[\frac{1}{1 + e^{-X_i B_2 - (\alpha_1 - \alpha_0) y_i + \alpha_1 B_{iH}}} \right]
\end{aligned}$$

가 (10) (12) α, β, B

(Maximum likelihood estimates: MLE)

3)

가

, 가 (controllability: CON), (extensiveness: EXT), (severity: SEV),
(familiarity: FAM), (dread: Dread), (risk: RISK),
가 (personal exposure: PER) 가 (public exposure: PUB) 8가

가

가

가

2

가

20%

가

100

1		20%	가
	5		2
		no-no	yes-yes
		가	
	2,500	1	

4.

		(expectation)	
		no-no	yes-yes
		가	
	, no-yes	yes-no	(B _{il} +B _i)/2
			(B _i +B _{iH})/2

Bishop and Heberlein

$$\begin{aligned}
 \overline{EWTP}_i &= E(\overline{WTP}_i) = 0 * \hat{P}_{i \in ZR} + (1 - \hat{P}_{i \in ZR}) E \left(\hat{\alpha}_i y_i + \frac{1}{\hat{\alpha}_1} \exp \left(\frac{X_i \hat{B}_2 - \varepsilon_i}{\hat{\beta}} \right) \right) \\
 &= \frac{1}{1 + e^{Z_i \hat{B}_1}} \left(\hat{\alpha}_i y_i + \frac{1}{\hat{\alpha}_1} \exp \left(\frac{X_i \hat{B}_2}{\hat{\beta}} \right) E \left(\exp \left(\frac{-\varepsilon_i}{\hat{\beta}} \right) \right) \right) \\
 &= \frac{1}{1 + e^{Z_i \hat{B}_1}} \left(\hat{\alpha}_i y_i + \frac{1}{\hat{\alpha}_1} \exp \left(\frac{X_i \hat{B}_2}{\hat{\beta}} \right) \frac{1}{\hat{\beta}} \Gamma \left(\frac{1}{\hat{\beta}} \right) \right)
 \end{aligned}$$

Cameron

$$\begin{aligned}
 \overline{EWTP}_i &= E(\overline{WTP}_i) = 0 * \hat{P}_{i \in ZR} + (1 - \hat{P}_{i \in ZR}) E \left(\frac{(\hat{\alpha}_1 - \hat{\alpha}_0)}{\hat{\alpha}_1} y_i + \frac{1}{\hat{\alpha}_1} X_i \hat{B}_2 + \frac{1}{\hat{\alpha}_1} \eta_i \right) \\
 &= \frac{1}{1 + e^{Z_i \hat{B}_1}} \left(\frac{(\hat{\alpha}_1 - \hat{\alpha}_0)}{\hat{\alpha}_1} y_i + \frac{1}{\hat{\alpha}_1} X_i \hat{B}_2 + \frac{1}{\hat{\alpha}_1} E(\eta_i) \right) \\
 &= \frac{1}{1 + e^{Z_i \hat{B}_1}} \left(\frac{(\hat{\alpha}_1 - \hat{\alpha}_0)}{\hat{\alpha}_1} y_i + \frac{1}{\hat{\alpha}_1} X_i \hat{B}_2 \right)
 \end{aligned}$$

	Root Mean Square Error (RMSE)	Mean Absolute Error (MAE)	2
2)	2	zero-response mixture model	Cameron
		zero-response	Cameron
		3	
	(Reduced Variables Model)		

$$WTP_i = \begin{cases} 0: P = \frac{e^{-0.816-0.214EXT_i-0.392EDC_i}}{1+e^{-0.816-0.214EXT_i-0.392EDC_i}} & (:) \\ 3.581+0.003y_i+0.310PUB_i-0.015AGE_i+1.52\eta_i:(1-P) & \end{cases}$$

$$WTP_i = \begin{cases} 0: P = \frac{e^{-0.237EXT_i+0.134FAM_i-0.126Dread_i-0.202RISK_i}}{1+e^{-0.237EXT_i+0.134FAM_i-0.126Dread_i-0.202RISK_i}} & (:) \\ 12-0.57CON_i+0.91Dread_i-0.082AGE_i+6.8\eta_i:(1-P) & \end{cases}$$

가

,

가

가

가

$$Mean = \frac{1}{N} \sum_{i=1}^N \overline{EWTP}_i = \frac{1}{N} \sum_{i=1}^N \frac{1}{1+e^{Z_i \hat{\beta}_1}} \left(\frac{(\hat{\alpha}_1 - \hat{\alpha}_0)}{\hat{\alpha}_1} y_i + \frac{1}{\hat{\alpha}_1} X_i \hat{B}_2 \right)$$

2)

$$RMSE = \sqrt{\frac{1}{N} \sum_{i=1}^N (\overline{EWTP}_i - SWTP_i)^2}, \quad MAE = \frac{1}{N} \sum_{i=1}^N |\overline{EWTP}_i - SWTP_i|$$

, 24,344 , 8.3
 가 , 20%
 100 4
 가 $(24,344)/(4/10^6)=60.8$ 가
 $8.3/((4/10^6)/2500)=51.9$ 가
 4
 2.5 , 1.3 가
 가
 “ ”
 가
 가
 가 (1995 3.3) 18.4 ,
 15.7 가
 186
 260 가
 가
 가
 11 가 2.8 165
 [6]. NRC 가 300 (2001 2
 1250 /\$ 37 5) (1999,[7])
 가 7 5 , 27 4

5.

가
 ,
 CVM 가 가
 가 가
 가
 가
 가
 가

[1] 4가 2가 , 가 가 가 (National Oceanic & Atmosphere Administration : NOAA) 가 가 가 'Scope Test' [8]. , 20% 20% 가 가 가 .

1. , , (2001), “가 가 가 : ”, 2001 ,
2. Bishop, R.C. and Heberlein, T.A. (1979), “Measuring values of extramarket goods: Are indirect measured biased?”, American J. of Agricultural Economics, 61(5), 926-930
3. Cameron, T.A. (1988) “A new paradigm for valuing non-market goods using referendum data. Maximum Likelihood estimation by censored logistic regression”, J. of Environmental Economics and Management, 15, 355-379
4. Kriström, Bengt (1995), “Spike Models in Contingent Valuation: Theory and Illustration”, Invited Paper to the 1st Toulouse Conference on Environmental and Resource Economics, Toulouse, France, March 30-31
5. An, Yuying and Roberto Ayala (1995), “A Mixture Model of Willingness to Pay Distribution”, Department of Economics Working Paper #95-21, Duke University
6. Baum, J.W. (1994) Value of Public Health and Safety Actions and Radiation Dose Avoided: NUREG/CR-6212
7. (1999), “ - (II)”, , KINS/HR-244
8. Arrow, K., Solow, R., Leamer, E., Portney, P., Radner, R., Schuman, H. (1993) “Report of the NOAA Panel on Contingent Valuation”, US Government Federal Register, January 15, Vol. 58 No. 10, pp. 4601-4614

1.

(,)
:
(500; 1,000; 10,000)
(1,000; 5,000; 10,000)
(1,000; 10,000; 20,000)
(10,000; 20,000; 50,000)
(20,000; 40,000; 80,000)

2. Bishop and Heberlein

Cameron

RMSE

MAE

	Bishop and Heberlein (zero-response)	Cameron (zero-response)
	RMSE=3.236 () MAE=2.627 ()	RMSE=2.927 () MAE=2.124 ()
	RMSE=15.59 () MAE=11.92 ()	RMSE=3.26 () MAE=2.01 ()

3.

Cameron

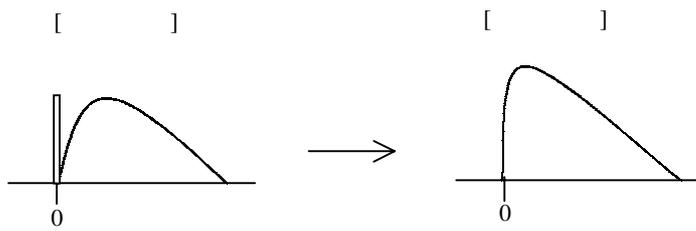
Parameters	Full Variables		Reduced Var.		Full Variables		Reduced Var.	
	t-value		t-value		t-value		t-value	
B₁								
INTERCEPT	-0.875	-1.283	-0.816	-2.944	0.457	0.633	-	-
Controllability	0	*	-	-	0.039	0.592	-	-
Severity	0	*	-	-	0	*	-	-
Extensiveness	-0.151	-1.794	-0.214	-2.904	-0.226	-2.367	-0.237	-3.153
Familiarity	0.012	0.141	-	-	0.145	2.452	0.134	2.368
Dread	-0.126	-0.419	-	-	-0.101	-1.093	-0.126	-1.510
Risk	0	*	-	-	-0.258	-3.454	-0.202	-4.036
Personal exposure	0	*	-	-	0	*	-	-
Public exposure	-0.081	-0.746	-	-	-0.006	-0.079	-	-
Age	-0.002	-0.311	-	-	0.002	0.355	-	-
Education level	-0.425	-4.867	-0.392	-5.475	-0.400	-4.403	-0.362	-6.328
B₂								
INTERCEPT	2.256	3.588	2.356	7.214	1.370	2.204	1.777	5.206
Controllability	-0.042	-0.736	-	-	-0.080	-1.401	-0.084	-1.503
Severity	0	*	-	-	0.068	0.640	-	-
Extensiveness	0	*	-	-	0.010	0.113	-	-
Familiarity	-0.048	-0.590	-	-	0	*	-	-
Dread	0.020	0.232	-	-	0.084	0.981	0.133	1.876
Risk	0	*	-	-	0	*	-	-
Personal exposure	0	*	-	-	0	*	-	-
Public exposure	0.221	2.234	0.204	2.314	0.004	0.057	-	-
Age	-0.012	-1.964	-0.01	-2.057	-0.007	-1.374	-0.012	-2.445
Education level	0.007	0.084	-	-	0.091	1.135	-	-
($\alpha_1 - \alpha_0$)	0.002	2.641	0.002	3.023	0.0002	0.394	-	-
α_1	0.658	26.62	0.659	26.7	1471.2	24.5	1468.7	24.5
Log(L)	-2461.2		-2463.8		-2417.2		-2419.3	
RMSE	2.927()		2.927()		3.26()		3.26()	
MAE	2.124()		2.126()		2.01()		2.01()	

<Note> * boundary solution .

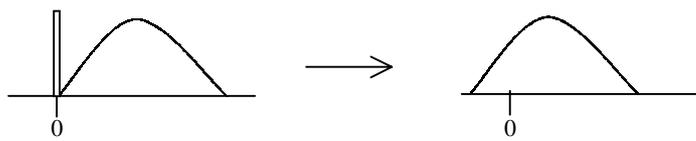
4. 가 (:)

1	24.1	60.8 (18.4)
1	39.3	51.9 (15.7)

() 가



(a) WTP가



(b) WTP가

1. Zero-response가

WTP

보 기 13

IV. 30~39세 여자

 2000년 상반기 통계에 의하면 30~39세 가구주는 월평균 228.9만원의 소득을 올렸으며 이 소득을 평균적으로 다음과 같은 항목에 지출하였습니다. 예를들어 평균적으로 보건 의료 분야에 6만원을 지출하였고 개인 교통 분야에는 14만원을 지출하였습니다.

<표 1> 30~39세 가구주 월평균 소비지출액

지 출 내 역	지 출 액
식 료 품	41 만원
의 식	17 만원
주 거	5 만원
광 열 수 도	9 만원
가 구 가 사	6 만원
피 복 신 발	8 만원
보 건 의 료	6 만원
교 육	14 만원
교 양 오락	9 만원
교 통 통 신	24 만원
개 인 교 통	14 만원
기 타 소비 지출	27 만원
잡 비	20 만원

 97년 통계에 의하면 35세의 여성이 40세에 이르지 못하고 사망할 확률은 '0.00502'이며 앞으로 더 살 수 있는 기간을 나타내는 평균 기대수명은 '44.59'년입니다. 다음은 특정사인에 의한 사망 관련자입니다. 예를들어 35세의 여성은 '뇌혈관질환'으로 사망할 확률이 '16.66%'로 가장 높으며 그 다음은 '심장질환', '당뇨병' 등의 순서입니다. 특히 '교통사고'에 의한 사망확률도 '1.74%'정도 됩니다.

<표 2> 35세 여성 100명은 남은 수명기간동안

16.66 명이 뇌혈관질환으로 사망하여 이 사인을 제거하면 2.02 년의 수명이 증가합니다.	
2.58 명이 고혈압성 질환으로	0.26 년의
7.41 명이 심장질환으로	0.83 년의
3.07 명이 위암으로	0.55 년의
1.57 명이 간암으로	0.32 년의
1.94 명이 폐암으로	0.33 년의
1.74 명이 교통사고로	0.42 년의
0.60 명이 자살로	0.16 년의
1.55 명이 간질환으로	0.31 년의
3.60 명이 당뇨병으로	0.55 년의
0.69 명이 결핵으로	0.10 년의
1.00 명이 폐렴으로	0.08 년의