# **Bayesian Model on Fatigue Crack Growth Rate of Type 304 Stainless Steel**



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### . Introduction

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The fatigue crack growth rate is typically estimated by deterministic methods in accordance with the ASME Boiler and Pressure Vessel Code Sec. XI

 $\rightarrow$  Deterministic model include uncertainties in the constant of the model

▷ Crack length was measured by direct current potential drop (DCPD) method using voltmeter and optical method using travelling microscope

 $\triangleright$  The relation for crack length and voltage is followed by Johnson's equation [4]

Where U is the potential drop (V),

▷ Likelihood function [5] - Assuming  $x_i - \dot{a}$  normal distribution is  $N(0, \sigma^2)$  $L(C,m \mid \dot{a}) = \prod^{N} \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{\sigma}}$ Where  $\sigma$  is the standard deviation of  $x_i - \dot{a}$ ▷ Prior and posterior C and m distributions

- Random sampled using Monte Carlo simulation Prior
Posterior (b)(a)

Uncertainty problems can be overcome by  $\triangleright$ probabilistic methods that estimate the degradation of materials even if the additional data scarcity of the fatigue model

The unknown constants of Paris' law were updated probabilistically by Bayesian inference  $\triangleright$  This methods can be used for the probabilistic structural integrity assessment of nuclear materials in the lab scale

# 2. Methods and Results

### **2.1 Materials**

▷ Polished Type 304 stainless steel was used

Table I. Chemical composition of Type 304 stainless steel

Chemical composition (%)	С	Si	Mn	Р	S	Ni	Cr	Мо	N	Cu	Fe	
Type 304 stainless steel	0.044	0.47	1.15	0.038	0.002	8.00	18.14	0.22	0.023	0.34	Bal.	(

Table II. Mechanical properties of Type 304 stainless steel (STS 304) measured at 25  $^{\circ}$ C in air condition in accordance



a is the crack length (mm),  $a_0$  is the reference crack length (mm), y is the length between notch centerline and the voltage measurement point shown in Fig. 1

 $\triangleright$  Constant load and constant  $\Delta K$  test conditions

Table III. Fatigue test conditions (left) and pre-cracking conditions (right)

Fatigue Test No.	1	2	
Mode	Const. load	Const. ΔK	
$\Delta \mathbf{P}(\mathbf{kN})$	18	-	
$\Delta \mathbf{K} \left( MPa\sqrt{m} \right)$	-	30	
R ratio (σ_min/σ_max)	0.1		
Frequency (Hz)	5		
Environment	Air		
Temperature (°C)	25		

### **2.3 Fatigue Test Results**



Const
load
18
0.1
10
Air
25
1

#### Fatigue Crack Growth Rate updated by Bayesian Inference Stainless Steel 304 Environment : Air Temperature : 25°C R (omin/omax) : 0.1 Frequency : 10 Hz (mm/cy 1E-4 da/dN da/dN by Constant Load da/dN by Constant delta K prior Paris' Law fitting posterior Paris' Law (updated by C) posterior Paris' Law (updated by m) dK (MPa\*m^0.5) Fig 7. Updated Paris' law results using Bayesian inference Table IV. Paris' law constant results of Type 304 stainless steel Sampled Posterior



Fig 6. Prior and posterior constant (a) C and (b) m distribution

with ASTM E8/E8M-15a (Straining rate=0.75mm/min)

Materials	STS 304
0.2 % offset yield strength (MPa)	264.4
Ultimate tensile strength (MPa)	757.2
Elastic modulus (GPa)	178.9
Elongation (%)	66.93

### **2.2 Fatigue Test Procedure**

- > The test was controlled by a servo-hydraulic testing control machine named Instron<sup>®</sup> Model 8516 with a load capacity of 100kN
- > SEN (Single Edge Notch) specimens were made by electrical-discharge machining (EDM) wire cutting
- $\triangleright$  The specimens were made in accordance with ASTM E647-13ae1, Gary S. Was et al., a dissertation from Il Soon Hwang, and a thesis from Jae Young Yoon [1-3]





#### Fig 2. Fatigue test picture (a) left side view (b) right side v



Fig 4. Picture of (a) notch

	Par cor	Prior	
	C	Mean	$1.5625 \times 10^{-10}$
	C	STD	5.8317 $\times$ 10 <sup>-12</sup>
	Par	Prior	
eei	COI	nstant	

Paris' law constant		Prior	Sampled prior	Posterior	
100	Mean	3.9388	3.9388	3.8834	
m	STD	$9.9560 \times 10^{-2}$	$9.9562 \times 10^{-2}$	$1.1213 \times 10^{-4}$	

prior

 $1.5626 \times 10^{-10}$ 

5.8337 ×  $10^{-11}$  | 2.1761 ×  $10^{-23}$ 

 $1.2923 \times 10^{-10}$ 

## **3.** Conclusions

▷ Paris' law constants C and m for Type 304 stainless steel were determined by probabilistic method using Bayesian inference

 $\rightarrow$  <u>Uncertainty of models' constant decreases dramatically</u>

▷ Until now, remaining lives of NPPs are estimated by deterministic methods using a priori model to finally assess structural integrity.

▷ Bayesian approach can utilize in-service data derived from aged properties

 $\rightarrow$  A probabilistic method should be applied to consider the environment and material conditions

#### Reference

Fig 1. SEN specimens for fatigue testing and wire attachment positions on specimens (unit : mm)

> Stress intensity factor of fatigue specimens  $K = \frac{P\sqrt{a}}{BW} \left| 1.986 + 1.782 \left(\frac{a}{W}\right) + 6.998 \left(\frac{a}{W}\right)^2 - 21.505 \left(\frac{a}{W}\right)^3 + 45.351 \left(\frac{a}{W}\right)^4 \right|$ 

until  $\frac{u}{W} = 0.621$  (it shows a maximum difference of 6%)

Where K is the stress intensity factor ( $MPa\sqrt{m}$ ), P is the applied load (N), a is the crack length (mm), B is the thickness of specimens (mm), and W is the width of specimens (mm)

by DCPD VS by microscope (b) crack with microscope

DCPD methods measured crack length very well [1] G. S. Was and R. G. Ballinger, "Hydrogen Induced Cracking

**2.4 Bayesian Updating** 

▷ Bayesian theorem [5] f(C,m): Prior distribution of constant C and m  $f(C, m \mid \dot{a})$ : Posterior distribution of constant C and m  $f(C, m | \dot{a}) = kL(C, m | \dot{a})f(C, m)$  $L(C, m | \dot{a})$ : Likelihood k: Normalizing constant ▷ Normal distribution [6]

#### - Probability density function (PDF) $(C, m-\mu_{C,r})$ $f(C,m) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{\sigma\sqrt{2\pi}}}}e^{-\frac{1}{\sigma\sqrt{2\pi}}}e^{-\frac{1}{\sigma\sqrt{2\pi}}$ $2\sigma_{C,m}^{2}$ Where $\mu$ is the mean of C and m,

 $\sigma$  is the standard deviation of C and m

DCPD a (mm)

Fig 5. Crack length measured

- Cumulative distribution function (CDF)

 $F(C,m) = \Phi\left(\frac{C,m-\mu_{C,m}}{\sigma_{C,m}}\right) = \frac{1}{2}\left|1 + \operatorname{erf}\left(\frac{C,m-\mu_{C,m}}{\sigma_{C,m}\sqrt{2}}\right)\right|$ Where  $\Phi(x) = \frac{1}{2} \left[ 1 + \operatorname{erf}\left(\frac{x}{\sqrt{2}}\right) \right], \operatorname{erf}(x) = \frac{1}{\sqrt{\pi}} \int_{-x}^{x} e^{-t^2} dt$ 

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